The Bayesian (nonparametric) approach to Statistics via exchangeability

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BAYSM 2016, Firenze (ITALY)

21st June 2016



Outline

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Introduction to Bayesian Statistics

Notation

🖙 The Bayesian paradigm and Bayes' Theorem

Read Exchangeability

Critical aspects of the Bayesian paradigm according to de Finetti

- Representation Theorem
- Re-interpretation of the Bayesian paradigm
- Parametric models

References

• □ > • □ > • □ >

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangeability

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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✓ Conditional probabilities and Bayes' Theorem provide a rational method for updating beliefs in the light of new information (DATA)

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Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of th Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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BAYESIAN INFERENCE

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BAYSM 2016, 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of th Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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• the prior belief is UPDATED via observed data and yields posterior distribution

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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BAYESIAN INFERENCE

✓ It is a typical scientific approach:

- the prior belief is UPDATED via observed data and yields posterior distribution
- it suggests that scientific inference is based on 2 parts: one depends on the scientist's subjective opinion and understanding of the phenomenon under study **BEFORE an EXPERIMENT** was performed, the other depends on the observed data the scientist has obtained from the experiment.

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References and notation

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Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

This lesson is based mainly on Regazzini (2015), an item of the encyclopedia Wiley StatsRef.

- X is a separable and complete metric space
- \mathscr{X} is the Borel σ -algebra of subsets of \mathbb{X}
- (X_n)_{n≥1} is a sequence of random elements defined on some probability space (Ω, ℱ, ℙ) and taking values in (X[∞], ℋ[∞])

 $(X_n)_{n\geq 1}$ is the sequence of observations (DATA); X_n is the result of the random experiment at trial n

References and notation

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A. Guglielmi

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BAYSM 2016, 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

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- $\mathbf{P}_{\mathbb{X}}$ space of all probability measures on $(\mathbb{X}, \mathscr{X})$, with the topology of weak convergence
- $\mathscr{P}_{\mathbb{X}}$ Borel σ -algebra of subsets of \mathbf{P}_X
- A random element p̃ defined on (Ω, 𝔅, ℙ) and taking values in (𝒫_𝔅, 𝒫_𝔅) is a random probability measure.

The Bayesian paradigm

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Introduction to Bayesian Statistics

Traditionally:

The Dave

paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

$X_1, X_2, X_3, \dots | \tilde{p} = p \stackrel{\text{i.i.d.}}{\sim} p$ "true" distribution of each observation $\tilde{p} \sim Q$ prior

$$\Rightarrow \pi(A_1 \times \cdots \times A_n \times B) := P(X_1 \in A_1, \dots, X_n \in A_n, \tilde{p} \in B) = \int_B \prod_{i=1}^n p(A_i)Q(dp)$$

 π is a probability on $\mathbb{X}^\infty\times \mathbf{P}_\mathbb{X}$

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The Bayesian paradigm

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Introduction to Bayesian Statistics

Traditionally:

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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 π is a probability on $\mathbb{X}^{\infty} \times \mathbf{P}_{\mathbb{X}}$

There exists a function $Q_n(B; x^{(n)}), B \in \mathscr{P}_{\mathbb{X}}, x^{(n)} := (x_1, \dots, x_n) \in \mathbb{X}^n$ such that $\checkmark B \mapsto Q_n(B; x^{(n)})$ is a probability on $\mathscr{P}_{\mathbb{X}}$ for all $x^{(n)}$; $\checkmark x^{(n)} \mapsto Q_n(B; x^{(n)})$ is \mathscr{X}^n -measurable for all B; $\checkmark Q_n(B; x^{(n)}) \stackrel{a.s.}{=} P(\tilde{p} \in B | X^{(n)})$ for all B $\Rightarrow Q_n(\cdot; X^{(n)})$ is the conditional distribution of \tilde{p} (the 'true" distribution of the data), given $X^{(n)} := (X_1, \dots, X_n)$ (\mathbb{X} Polish space)

The Bayesian paradigm

Introduction to Bayesian Statistics

Traditionally:

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpre tation of the Bayesian paradigm

Parametric models

References

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 $Q_n(\cdot; X^{(n)})$ is the **posterior distribution** of \tilde{p} , given observations X_1, \ldots, X_n

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Bayes' Theorem for dominated models

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Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

 $X_i | \tilde{\theta} = \theta \stackrel{\text{i.i.d.}}{\sim} f_{\theta}(\cdot) \text{ "true" density of each observation} \\ \tilde{\theta} \sim \pi \text{ probability on } \Theta \text{ Euclidean space}$

Interpretation: $\theta \mapsto \prod_{i=1}^{n} f_{\theta}(x_i)$ likelihood, $\pi(d\theta)$ prior

Bayes' Theorem for dominated models

A. Guglielmi

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Interpretation: $\theta \mapsto \prod_{i=1}^{n} f_{\theta}(x_i)$ likelihood, $\pi(d\theta)$ prior

Then the posterior distribution of $\tilde{\theta}$, given $X_1 = x_1, \dots, X_n = x_n$, can be computed by **Bayes' Theorem**:

$$P(\tilde{\theta} \in B | X_1 = x_1, \dots, X_n = x_n) \stackrel{a.s.}{=} \frac{\int_B \prod_{i=1}^n f_\theta(x_i) \pi(d\theta)}{\int_\Theta \prod_{i=1}^n f_\theta(x_i) \pi(d\theta)}, \quad B \in \mathscr{B}(\Theta)$$

Proof: definition of conditional distribution (as the solution of an integral equation) + Radon-Nikodym Theorem

The prior distribution

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Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Making inference about $\tilde{\theta}$ is to learn about the unknown $\tilde{\theta}$ from the **DATA**: based on the data, explore which values of $\tilde{\theta}$ are probable, what might be plausible numbers as estimates of the different components of θ and the extent of uncertainty associated with such estimates

the distribution of $\tilde{\theta}$, **prior distribution**: it quantifies the uncertainty about $\tilde{\theta}$ prior to seeing data

The prior distribution

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BAYSM 2016. 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of th Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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The prior represents the subjective belief and knowledge \mapsto subjective prior,

or, a conventional prior supposed to represent small or no information \mapsto noninformative/vague/objective prior

The prior distribution

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of th Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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or, a conventional prior supposed to represent small or no information \mapsto noninformative/vague/objective prior **\frac{1}{2}** ARGH!!!! **\frac{1}{2}** Jupiter Fulminator could hurl thunderbolts to us at any moment since now, since we follow, at least in principle,

Bruno de Finetti's approach to Bayesian Statistics, i.e. the subjective approach!

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BAYSM 2016, 21/06/16

Introduction to Bayesian Statistics

- Notation
- The Bayesian paradigm and Bayes' Theorem
- Exchangeability
- Critical aspects of the Bayesian paradigm according to de Finetti
- Representatio Theorem
- Re-interpretation of the Bayesian paradigm
- Parametric models
- References

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Bruno de Finetti

Introduction to Bayesian Statistics

- Notation
- The Bayesian paradigm and Bayes' Theorem
- Exchangeability
- Critical aspects of the Bayesian paradigm according to de Finetti
- Representation Theorem
- Re-interpretation of the Bayesian paradigm
- Parametric models
- References

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A. Guglielmi

BAYSM 2016, 21/06/16

Bruno de Finetti

Introduction to Bayesian Statistics

- Notation
- The Bayesian paradigm and Bayes' Theorem
- Exchangeability
- Critical aspects of the Bayesian paradigm according to de Finetti
- Representation Theorem
- Re-interpretation of the Bayesian paradigm
- Parametric models
- References

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• 1935: lectures at the Institut Poincaré, Paris (France); de Finetti (1937)

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Definition. The sequence $(X_n)_{n\geq 1}$ is exchangeable if

 $(X_1,\ldots,X_n)\stackrel{\mathrm{d}}{=}(X_{\pi(1)},\ldots,X_{\pi(n)})$

for any $n \ge 1$ and permutation π of $(1, \ldots, n)$.

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A. Guglielmi

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Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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• According to this definition, the order with which data are recorded is irrelevant for inferential purposes

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Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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- it is a *weak* assumption, and translates lack of (enough) information through a condition of *symmetry*

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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- According to this definition, the order with which data are recorded is irrelevant for inferential purposes
- it is a *weak* assumption, and translates lack of (enough) information through a condition of *symmetry*
- · For example, in coin-tossing sequence one would have

$$\mathbb{P}[X_1 = 0, X_2 = 1, X_3 = 1, X_4 = 0] = \mathbb{P}[X_1 = 1, X_2 = 0, X_3 = 0, X_4 = 1]$$

Critical aspects of the Bayesian paradigm according to de Finetti (Regazzini, 2015)

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

De Finetti gave his first results on exchangeability (he used the term equivalence) for sequences of trials on a given phenomenon, all made under analogous conditions, i.e. $(X_n)_n 0$ -1 r.v.'s, $X_n = 1$ if the *n*-th trial was a "success"

A. Guglielmi BAYSM 2016, 21/06/16

A 3 b

Critical aspects of the Bayesian paradigm according to de Finetti (Regazzini, 2015)

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangeability

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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✓ Phenomena whose trials are independent with a fixed but unknown probability distribution (p.d.) are exchangeable, since the law of $(X_n)_n$ is a mixture of Bernoulli r.v.'s.

A. Guglielmi BAYSM 2016, 21/06/16

4 3

Critical aspects of the Bayesian paradigm according to de Finetti (Regazzini, 2015)

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangeability

Critical aspects of the Bayesian paradigm according to de Finetti

Representati Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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✓ Phenomena whose trials are independent with a fixed but unknown probability distribution (p.d.) are exchangeable, since the law of $(X_n)_n$ is a mixture of Bernoulli r.v.'s.

✓ This statement is controversial according to de Finetti's subjectivist conception of probability; the reference to an unknown probability is devoid of sense and, in any case, obscure and specious. In addition, it requires the specification of a law for the fixed, though unknown, p.d., and, under the subjective point of view, such request is of an ambiguous content.

In more modern terms: if $X_i | \theta \stackrel{\text{i.i.d.}}{\sim} Be(\theta), \theta \sim F$ and the approach taken is subjective, then

 $\theta \sim F$ is ambiguous

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A. Guglielmi

BAYSM 2016. 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

 \checkmark *F*, the unknown p.d. of θ , cannot represent the subjective belief unless θ has an objective (i.e. "physical") meaning, that is, unless θ represents a well-specified characteristic of the members of a statistical population.

A. Guglielmi BAYSM 2016, 21/06/16

A 3 b

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

 \checkmark *F*, the unknown p.d. of θ , cannot represent the subjective belief unless θ has an objective (i.e. "physical") meaning, that is, unless θ represents a well-specified characteristic of the members of a statistical population.

Two examples:

- the sequence of drawings with replacement from an urn containing white and nonwhite balls according to an unknown composition: X_n = 1 if the *n*-th drawn ball is white, θ is the unknown proportion of white ball.
- the sequence of tossing of the same coin: $X_n = 1$ if the *n*-th toss is H, θ is the unknown probability of H in each toss.

A. Guglielmi BAYSM 2016, 21/06/16

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Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Two examples:

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 The composition is empirically verifiable!
- the sequence of tossing of the same coin: $X_n = 1$ if the *n*-th toss is H, θ is the unknown probability of H in each toss. The probability of H cannot be verified!

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BAYSM 2016. 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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✓Unlike these formulations, which are unfortunate and unclear in some respect, the condition of exchangeability they met is more general, always meaningful and sensible. De Finetti selected it to indicate a sufficiently general setting in which he might prove the validity of the principle of induction:

in a sequence of homogeneous trials, the frequency distribution of the results of n past trials can represent a good approximation, if n is "large", to the conditional distribution of the outcome of a future trial, given the observed frequency distribution.

A. Guglielmi

BAYSM 2016. 21/06/16

de Finetti's representation theorem for events

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Theorem. The sequence $X^{(\infty)} = (X_n)_{n\geq 1}$ of 0-1 r.v.s's is exchangeable if and only if there exists a probability measure F on $([0, 1], \mathscr{B}([0, 1]))$ such that

$$P[X_1 = x_1, \dots, X_n = x_n] = \int_{[0,1]} \theta^{\sum_{i=1}^n x_i} (1-\theta)^{n-\sum_{i=1}^n x_i} F(\mathrm{d}\theta)$$

for any $n \ge 1$ and $(x_1, ..., x_n)$ in $\{0, 1\}^n$.

A. Guglielmi BAYSM 2016, 21/06/16

de Finetti's representation theorem for events

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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for any
$$n \ge 1$$
 and $(x_1, ..., x_n)$ in $\{0, 1\}^n$.

Moreover, when $(X_n)_{n\geq 1}$ is exchangeable (since *P* is σ -additive):

•
$$\frac{\sum_{i=1}^{n} X_i}{n} \xrightarrow{a.s.} \tilde{\theta} \sim F \text{ as } n \to +\infty.$$

• Conditionally on
$$\tilde{\theta}$$
, $X_1, \ldots, X_n | \tilde{\theta} \stackrel{\text{i.i.d.}}{\sim} Be(\tilde{\theta})$ for all n
 $\tilde{\theta} \sim F$.

A. Guglielmi BAYSM 2016, 21/06/16

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Reinterpretation of the Bayesian paradigm through exchangeability

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

✓ Therefore, there is a formal equivalence between exchangeable trials of the same phenomenon (i.e. X_n 's are binary) and those trials that are designated as "independent, with a fixed, but unknown, probability".

A. Guglielmi BAYSM 2016, 21/06/16

Reinterpretation of the Bayesian paradigm through exchangeability

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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 \checkmark The representation theorem may be used to justify the principle of induction with reference to an excheangeable sequence of events:

If $(X_n)_n$ are exchangeable 0-1 r.v.'s, and $\varphi_n := \sum_{i=1}^n X_i/n$, then

$$\left|P\left(X_{n+1}=e_1,\ldots,X_{n+k}=e_k|X_1,\ldots,X_n\right)-\varphi_n^{\sum_{i=1}^k e_i}(1-\varphi_n)^{k-\sum_{i=1}^k e_i}\right| \stackrel{\text{a.s.}}{\to} 0, n \to +\infty,$$

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A. Guglielmi

BAYSM 2016, 21/06/16

for all $(e_1, \ldots, e_k) \in \{0, 1\}^k$, $k = 1, 2, 3, \ldots$.

Reinterpretation of the Bayesian paradigm through exchangeability

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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A. Guglielmi

BAYSM 2016, 21/06/16

for all
$$(e_1, \ldots, e_k) \in \{0, 1\}^k$$
, $k = 1, 2, 3, \ldots$.

✓ Roughly speaking its says that:

the frequency of success in n past trials can be used to evaluate the probability distribution of "future" trials of the same phenomenon, if n is large.

Comments on de Finetti's work on exchangeability

Introduction to Bayesian Statistics

- Notation
- The Bayesian paradigm and Bayes' Theorem
- Exchangeability
- Critical aspects of the Bayesian paradigm according to de Finetti
- Representatio Theorem
- Re-interpretation of the Bayesian paradigm
- Parametric models
- References

- ✓ De Finetti's re-interpretation of the Bayesian approach is strictly linked to the subjective notion of probability (i.e. definition of probability of an event through coherence): the probability *p* of an event *E* is the personal degree of belief in the event; *p* has to ensure that there exists no real value *c* such that any bet on *E* with gain $c(p \mathbf{1}_E)$ has realizations of the gain all strictly positive or all strictly negative
- ✓ De Finetti's original approach to exchangeability was different from what I have introduced here; see Cifarelli and Regazzini (1996) for an historical picture of de Finetti's contributions
- exchangeability is a term introduced by Pólya; de Finetti used *eventi equivalenti* (equivalent events), other authors used symmetric
- ✓ the infinite exchangeability is a keypoint here: finite exchangeable sequences have different representations.

de Finetti's representation theorem (general case) (1933)

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of th Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Theorem. The sequence $X^{(\infty)} = (X_n)_{n\geq 1}$ is exchangeable if and only if there exists a probability measure on $(\mathbf{P}_{\mathbb{X}}, \mathscr{P}_{\mathbb{X}})$ such that

$$P[X_1 \in A_1, \ldots, X_n \in A_n] = \int_{\mathbf{P}_{\mathbb{X}}} \prod_{i=1}^n p(A_i) Q(\mathrm{d}p)$$

for all $n \ge 1$ and A_1, \ldots, A_n in \mathscr{X} , where the probability Q is uniquely determined.

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de Finetti's representation theorem (general case) (1933)

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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for all $n \ge 1$ and A_1, \ldots, A_n in \mathscr{X} , where the probability Q is uniquely determined.

⇒ $(X_n)_{n\geq 1}$ is exchangeable if and only if there exists a random probability measure \tilde{p} on $(\mathbb{X}, \mathscr{X})$ such that $\tilde{p} \sim Q$ and

$$\mathbb{P}\left[X_1 \in A_1, \ldots, X_n \in A_n \,|\, \tilde{p}\,\right] = \prod_{i=1}^n \tilde{p}(A_i)$$

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A. Guglielmi

BAYSM 2016, 21/06/16

for any $n \ge 1$ and A_1, \ldots, A_n in \mathscr{X} .

de Finetti's representation theorem (general case) (1933) - continued

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Q is a probability measure on $\mathbf{P}_{\mathbb{X}} \longrightarrow de$ *Finetti measure* of $(X_n)_{n\geq 1}$

 \square If $(X_n)_{n\geq 1}$ is exchangeable, then its empirical distribution is such that

$$\frac{1}{n}\sum_{i=1}^{n}\delta_{X_{i}} \Rightarrow \tilde{p} \qquad \text{a.s.-}\mathbb{P}$$

where \Rightarrow denotes *weak convergence*.

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de Finetti's representation theorem (general case) (1933) - continued

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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where \Rightarrow denotes *weak convergence*.

Hierarchical representation: $(X_n)_{n\geq 1}$ exchangeable is equivalent to

 $X_i | \tilde{p} \stackrel{\text{i.i.d.}}{\sim} \tilde{p}$ $\tilde{p} \sim Q$ Q = prior distribution

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de Finetti's representation theorem (general case) (1933) - continued

Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representati Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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ⁿ Hierarchical representation: $(X_n)_{n\geq 1}$ exchangeable is equivalent to

 $\begin{aligned} X_i \mid & \tilde{p} \quad \stackrel{\text{i.i.d.}}{\sim} \quad & \tilde{p} \\ & & \tilde{p} \quad \sim \quad Q \\ & & Q \quad = \quad \text{prior distribution} \end{aligned}$

The Bayesian nonparametric framework is equivalent to exchangeability of $(X_n)_n$

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BAYSM 2016, 21/06/16

Parametric case through the representation theorem

A. Guglielmi

-

BAYSM 2016, 21/06/16

to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Parametric model: *Q* degenerate on a finite–dimensional subset $\mathbf{P}_{\mathbb{X}}^*$ of $\mathbf{P}_{\mathbb{X}}$, such that

$$Q(\{\mathbf{P}_{\mathbb{X}}^*\}) = Q(\{p \in \mathbf{P}_{\mathbb{X}} : p = p_{\theta}, \theta \in \Theta\}) = 1$$

and and there exists a function

$$\tilde{\theta}: \mathbf{P}^*_{\mathbb{X}} \to \Theta$$
 bijective.

 $\Theta \subset \mathbb{R}^p$ is called parametric space. The prior Q induces a probability on Θ :

$$\pi(B) = Q(\tilde{\theta}^{-1}(B)), B \in \mathscr{B}(\Theta).$$

Parametric case through the representation theorem

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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 bijective.

 $\Theta \subset \mathbb{R}^p$ is called parametric space. The prior Q induces a probability on Θ :

$$\pi(B) = Q(\tilde{\theta}^{-1}(B)), B \in \mathscr{B}(\Theta).$$

In these cases:

$$X_i \mid \tilde{\theta} = \theta \stackrel{\text{i.i.d.}}{\sim} p_{\theta}$$

 $\tilde{ heta} \sim \pi$ prior distribution

For instance:

$$Q(\{p \in \mathbf{P}_{\mathbb{X}} : p(\mathrm{d}x) = \varphi((x-\mu)/\sigma) \,\mathrm{d}x, \, (\mu,\sigma) \in \mathbb{R} \times \mathbb{R}^+\}) = 1$$

with φ being the density function of a N(0, 1) distribution.

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Introduction to Bayesian Statistics

Notation

The Bayesia paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

When can we assume $Q({\mathbf{P}_{\mathbb{X}}^*}) = 1$, where $\mathbf{P}_{\mathbb{X}}^*$ is finite-dimensional? More clearly, when could we assume the model is parametric?

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 BAYSM 2016, 21/06/16

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Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representatio Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

When can we assume $Q({\mathbf{P}_{\mathbb{X}}^*}) = 1$, where $\mathbf{P}_{\mathbb{X}}^*$ is finite-dimensional? More clearly, when could we assume the model is parametric?

• if, from past experience in cases similar to the one analized, we believe that the parametric family approximates well the "true" distribution

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BAYSM 2016. 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of th Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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- if, from past experience in cases similar to the one analized, we believe that the parametric family approximates well the "true" distribution
- if, in addition to exchangeability, we assume different conditions for the sequence of observations. For example, if (X_n)_{n≥1} is also spherically symmetric (L(X₁,...,X_n)^T = L(A(X₁,...,X_n)^T) for any orthogonal matrix A), then P^{*}_X is the family of Gaussian distributions with 0-mean.

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A. Guglielmi

BAYSM 2016. 21/06/16

Introduction to Bayesian Statistics

Notation

The Bayesiar paradigm and Bayes' Theorem

Exchangea bility

Critical aspects of the Bayesian paradigm according to de Finetti

Representati Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Otherwise: nonparametric model

 \longrightarrow greater flexibility when Q has large support, possibly supp $(Q) = \mathbf{P}_{\mathbb{X}}$.

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A. Guglielmi

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangeability

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

Many thanks to:

- \checkmark Eugenio Regazzini, who suggested the material I used
- \checkmark Antonio Lijoi for providing the slides with the notation and representation theorems

Acknowledgments

Introduction to Bayesian Statistics

Notation

The Bayesian paradigm and Bayes' Theorem

Exchangeability

Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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Bibliography

Introduction to Bayesian Statistics

Notation

- The Bayesia paradigm and Bayes' Theorem
- Exchangea bility
- Critical aspects of the Bayesian paradigm according to de Finetti

Representation Theorem

Re-interpretation of the Bayesian paradigm

Parametric models

References

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A. Guglielmi

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